Open-Set Recognition with Gaussian Mixture Variational Autoencoders: Supplementary Material

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1 Neural network assumptions

We call a neural network f_{τ} an *n*-headed neural network if

- 1. $f_{\tau}: \mathbb{R}^m \to \prod_{i=1}^n \mathbb{R}^s$, i.e. it maps b to $(a_1, a_2, ..., a_n)$ with $a_i \in \mathbb{R}^s$,
- 2. for each $i, 1 \le i \le n$, we have $a_i = f_{\ell_i}^i \circ f_{\ell_i-1}^i \circ \ldots \circ f_{t+1}^i \circ f_t \circ f_{t-1} \circ \ldots \circ f_1(b)$ for an integer t not depending on $i, \ell_i \ge t+1$, and each f_j, f_j^i is a typical neural network single layer parameterized by a matrix and a bias vector, and it includes an activation function. Vector τ corresponds to all these parameters.

In GMVAE, neural networks corresponding to q_{ϕ_z}, q_{ϕ_w} are 2-headed neural networks (mean and covariance) with ϕ_z, ϕ_w denoting all of the respective parameters. Probability p_{θ} is a 1 or 2-headed network with parameters θ , and p_{β} for $\beta = (\beta_{K_1}, \beta_{K_2}, ..., \beta_{K_C})$ consists of a $(2 \sum_{c=1}^{C} K_c)$ -headed neural network.

Assumption 1. In each network q_{ϕ_z} , q_{ϕ_w} , p_{θ} , and p_{β} , the last layer in each head $f_{\ell_i}^i$ has an identity activation function.

Assumption 2. Neural network $p_{\beta'}$ for $\beta' = (\beta_{K_1}, ..., \beta_{K_c+1}, ..., \beta_{K_C})$ consists of p_{β} with simply two additional heads, while all other network architectures are the same.

Lemma 1. Under Assumption 1 for an n-headed network, we have that given any $\overline{a} = (\overline{a}_1, ..., \overline{a}_n)$, there exists $\tau = \tau(\overline{a})$ such that $f_{\tau}(b) = \overline{a}$ for every b.

Proof. Let \overline{a} be given. We define τ to consist of 0 matrices and biases for each layer except $f_{\ell_i}^i$. In $f_{\ell_i}^i$, the matrix is 0 but the bias is \overline{a}_i . Since $f_{\ell_i}^i$ has the identity activation, it follows $f_{\tau}(b) = \overline{a}$ for every b.

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2 **Proof of Proposition 1**

Proposition 1. Let us assume that $x \in \mathcal{X}$ is distributed as $x \sim p_{data} = \mathcal{B}(\mu_x)$, C = 1, and Assumption 1 holds. Then the optimal GMVAE loss is constant with respect to K. In fact, we have that $\min -\mathbb{E}_{\mathcal{X}}[\mathcal{L}(K)] = -\mathbb{E}_{\mathcal{X}}[\log p_{data}]$ for every $K \ge 1$ and a globally optimal solution reads

$$\begin{array}{ll}
\mu(x;\phi_{z}^{*}) &= \mu_{c=1,k}(w;\beta^{*}) = \mu_{z} \\
\sigma^{2}(x;\phi_{z}^{*}) &= \sigma_{c=1,k}^{2}(w;\beta^{*}) = \sigma_{z}^{2} \\
\mu(x,y;\phi_{w}^{*}) &= \vec{0} \\
\sigma^{2}(x,y;\phi_{w}^{*}) &= \vec{1} \\
\mu(z;\theta^{*}) &= \mu_{x}
\end{array}$$
(1)

for any constant vectors μ_z, σ_z .

Proof. Note that $(\phi_z^*, \phi_w^*, \beta^*, \theta^*)$ exist due to Assumption 1 and Lemma 1. First, we show that (θ^*, β^*) given in (1) maximize the log likelihood $\mathbb{E}_{\mathcal{X}} [\log p_{\theta,\beta}(x|y=1)]$ and results in $p_{\theta^*,\beta^*}(x|y=1) = p_{\text{data}}$. We have

$$KL(p_{\text{data}}||p_{\theta,\beta}(x|y=1)) = \mathbb{E}_{\mathcal{X}}\left[\log p_{\text{data}}\right] - \mathbb{E}_{\mathcal{X}}\left[\log p_{\theta,\beta}(x|y=1)\right]$$

and thus maximizing $\mathbb{E}_{\mathcal{X}} [\log p_{\theta,\beta}(x|y=1)]$ is equivalent to minimizing $KL(p_{\text{data}}||p_{\theta,\beta}(x|y=1))$. The global minimum of $KL(p_{\text{data}}||p_{\theta,\beta}(x|y=1))$ is clearly when $p_{\text{data}} = p_{\theta,\beta}(x|y=1)$. This is indeed the case for (θ^*, β^*) , since

$$p_{\theta^*,\beta^*}(x|y=1) = \int_{w,z,v} p_{\beta^*,\theta^*}(x,v,w,z|y=1) dw dz dv$$

=
$$\int_{w,z,v} p_{\theta^*}(x|z) p_{\beta^*}(z|w,y=1,v) p(v|y=1) p(w) dw dz dv$$

=
$$\int_{w,z,v} p_{\text{data}} p_{\beta^*}(z|w,y=1,v) p(v|y=1) p(w) dw dz dv$$

=
$$p_{\text{data}}$$
 (2)

because of GMVAE's generative model factorization and (1). Now we have

$$\mathbb{E}_{\mathcal{X}}\left[\log p_{\text{data}}\right] = \mathbb{E}_{\mathcal{X}}\left[\log p_{\theta^*,\beta^*}(x|y=1)\right]$$

$$= \mathbb{E}_{\mathcal{X}}\left[\mathbb{E}_{q_{\phi^*}(v,w,z|x,y=1)}\left[\log \frac{p_{\theta^*,\beta^*}(x,z,w,v|y=1)}{q_{\phi^*}(v,w,z|x,y=1)}\right]\right]$$

$$+ \mathbb{E}_{\mathcal{X}}\left[\mathbb{E}_{q_{\phi^*}(v,w,z|x,y=1)}\left[\log \frac{q_{\phi^*}(v,w,z|x,y=1)}{p_{\theta^*,\beta^*}(z,w,v|x,y=1)}\right]\right]$$
(3)

$$= \mathbb{E}_{\mathcal{X}}[\mathcal{L}(K;\phi_z^*,\phi_w^*,\beta^*,\theta^*)] + \mathbb{E}_{\mathcal{X}}[\mathrm{VG}(\phi_z^*,\phi_w^*,\beta^*,\theta^*)]$$
(4)

where $VG(\phi_z^*, \phi_w^*, \beta^*, \theta^*)$ corresponds to (3). We next show that $VG(\phi_z^*, \phi_w^*, \beta^*, \theta^*) = 0$. This together with the facts that maximized $\mathbb{E}_{\mathcal{X}}[\mathcal{L}(K; \phi_z, \phi_w, \beta, \theta)]$ corresponds with minimized $\mathbb{E}_{\mathcal{X}}[VG(\phi_z, \phi_w, \beta, \theta)]$, and $VG(\phi_z, \phi_w, \beta, \theta) \ge 0$ (it is a KL divergence), shows optimality.

From (1) we have that $p_{\theta^*}(x|z) = p_{\text{data}}(x)$ for all x and z and thus with (2) we have

$$p_{\theta^*,\beta^*}(z,w,v|x,y=1) = \frac{p_{\theta^*}(x|z,w,v,y=1)p_{\beta^*}(z,w,v|y=1)}{p_{\theta^*,\beta^*}(x|y=1)}$$
$$= \frac{p_{\theta^*}(x|z)p_{\beta^*}(z,w,v|y=1)}{p_{\text{data}}(x)}$$
$$= p_{\beta^*}(z,w,v|y=1).$$
(5)

The reconstruction term $p_{\theta}(x|z, w, v, y = 1) = p_{\theta}(x|z)$ for every θ because in GMVAE, data reconstruction depends only on z and is independent of w and v (see §3.1 of the paper).

Also from Bayes' and GMVAE's generative model factorization, we have the following simplification

$$p_{\beta^*}(v|z, w, y = 1) = \frac{p_{\beta^*}(z|w, y = 1, v)p(v|y = 1)p(w)}{p_{\beta^*}(z, w|y = 1)}$$

$$= \frac{p_{\beta^*}(z|w, y = 1, v)p(v|y = 1)p(w)}{p_{\beta^*}(z|w, y = 1)p(w|y = 1)}$$

$$= \frac{p_{\beta^*}(z|w, y = 1, v)p(v|y = 1)}{\sum_{v'} p_{\beta^*}(z|w, y = 1, v')p(v'|y = 1)}$$

$$= p(v|y = 1)$$
 (6)

where (1) is only used in the last line. Substituting (5) into $VG(\phi_z^*, \phi_w^*, \beta^*, \theta^*)$ we obtain $VG(\phi_z^*, \phi_w^*, \beta^*, \theta^*)$

$$\begin{split} &= \mathbb{E}_{q_{\phi^*}(v,w,z|x,y=1)} \left[\log \frac{q_{\phi^*}(v,w,z|x,y=1)}{p_{\theta^*,\beta^*}(z,w,v|x,y=1)} \right] \\ &= \mathbb{E}_{q_{\phi^*}(v,w,z|x,y=1)} \left[\log \frac{q_{\phi^*}(v,w,z|x,y=1)}{p_{\beta^*}(z,w,v|y=1)} \right] \\ &= \mathbb{E}_{p_{\beta^*}(v|z,w,y=1)q_{\phi^*_w}(w|x,y=1)q_{\phi^*_z}(z|x)} \left[\log \frac{p_{\beta^*}(v|z,w,y=1)q_{\phi^*_w}(w|x,y=1)q_{\phi^*_z}(z|x)}{p_{\beta^*}(z|w,y=1,v)p(w)p(v|y=1)} \right] \\ &= \mathbb{E}_{q_{\phi^*_w}(w|x,y=1)q_{\phi^*_z}(z|x)} \left[\log q_{\phi^*_z}(z|x) - \sum_{j=1}^K p_{\beta^*}(v=j|z,w,y=1) \log p_{\beta^*}(z|w,y=1,v=j) \right] \end{split}$$

$$+ KL(q_{\phi_w^*}(w|x, y = 1)||p(w)) \\+ \mathbb{E}_{q_{\phi_w^*}(w|x, y = 1)q_{\phi_z^*}(z|x)} [KL(p_{\beta^*}(v|z, w, y = 1))||p(v|y = 1))] \\= 0$$

due to (1) and (7). To complete the proof, simply note that negating (4) yields $-\mathbb{E}_{\mathcal{X}}[\mathcal{L}(K; \phi_z^*, \phi_w^*, \beta^*, \theta^*)] = -\mathbb{E}_{\mathcal{X}}[\log p_{\text{data}}].$

3 Proof of Proposition 2

Lemma 2. For every $\delta > 0$ and μ , there exists σ^2 such that if f(z) is the pdf of a d-dimensional Normal random vector with mean μ and diagonal covariance σ^2 then

$$f(z) \leq \delta$$
 for every z.

Proof. Let $u = \left(\frac{1}{\delta}(2\pi)^{-d/2}\right)^{1/d}$ and $\sigma = (u, ..., u)$. We have

$$f(z) = \prod_{i} \frac{1}{\sigma_i \sqrt{2\pi}} \exp\left\{-\frac{1}{2\sigma_i^2} (z_i - \mu_i)^2\right\} \le \prod_{i} \frac{1}{\sigma_i \sqrt{2\pi}} = \delta.$$

Proposition 2. Let us assume C = 1, Assumptions 1 and 2 hold, and that p(v|y = 1) is uniform in the appropriate dimension. We have

$$\min\left\{-\mathbb{E}_{\mathcal{X}}[\mathcal{L}(K;\phi_{z},\phi_{w},\beta,\theta)]\right\}-\min\left\{-\mathbb{E}_{\mathcal{X}}[\mathcal{L}(K+1;\phi_{z},\phi_{w},\beta,\theta)]\right\}\geq\epsilon_{K}$$

where $-\log 2 \le \log(K/(K+1)) \le \epsilon_K$ for all K.

Proof. We show that for every solution $(\phi'_z, \phi'_w, \beta', \theta')$ to min $\mathbb{E}_{\mathcal{X}}[-\mathcal{L}(K; \phi_z, \phi_w, \beta, \theta)]$, there exists a corresponding solution $(\phi^*_z, \phi^*_w, \beta^*, \theta^*)$ such that

$$-\mathbb{E}_{\mathcal{X}}[\mathcal{L}(K;\phi'_{z},\phi'_{w},\beta',\theta')] = -\mathbb{E}_{\mathcal{X}}[\mathcal{L}(K+1;\phi^{*}_{z},\phi^{*}_{w},\beta^{*},\theta^{*})] + \epsilon_{K}$$

Let us assume that $(\phi'_z, \phi'_w, \beta', \theta')$ minimizes $-\mathbb{E}_{\mathcal{X}}[\mathcal{L}(K; \phi_z, \phi_w, \beta, \theta)]$. Then we can choose

$$\begin{aligned}
\phi_z^* &= \phi_z' \\
\phi_w^* &= \phi_w' \\
\theta^* &= \theta'
\end{aligned}$$
(8)

which is a valid choice by Assumption 2, and have β^* such that

$$p_{\beta^*}(z|w, y = 1, v) = p_{\beta'}(z|w, y = 1, v) \quad \text{for all } v \le K$$
(9)

$$p_{\beta^*}(z|w, y = 1, v = K + 1) \le \delta \quad \text{for every } z, w \tag{10}$$

for any fixed $0 < \delta < 1/e$. Conditions (9) and (10) are always possible due to Assumptions 1 and 2 and Lemmas 1 and 2. In essence, we choose β^* such that the first K subcluster generative distributions are the same as the case β' but we take the (K + 1)-th subcluster generative distribution to map all points w to the same Normal distribution with large enough covariance.

Inserting (9) and (10) into (6) and combined with uniform priors, we get that

$$p_{\beta^*}(v = K + 1 | z, w, y = 1) = \frac{p_{\beta^*}(z | w, y = 1, v = K + 1)}{\sum_{j=1}^K p_{\beta'}(z | w, y = 1, v = j) + p_{\beta^*}(z | w, y = 1, v = K + 1)}$$
(11)

and

$$p_{\beta^*}(v=k|z,w,y=1) = \frac{p_{\beta'}(z|w,y=1,v=k)}{\sum_{j=1}^{K} p_{\beta'}(z|w,y=1,v=j) + p_{\beta^*}(z|w,y=1,v=K+1)} \\ \leq \frac{p_{\beta'}(z|w,y=1,v=k)}{\sum_{j=1}^{K} p_{\beta'}(z|w,y=1,v=j)} = p_{\beta'}(v=k|z,w,y=1)$$
(12)

for all $k \leq K$. The absolute difference between the two posteriors for $k \leq K$ in (12) is bounded by a factor of δ as follows:

$$\begin{aligned} \left| p_{\beta^*}(v=k|z,w,y=1) - p_{\beta'}(v=k|z,w,y=1) \right| \\ &= \left| \frac{p_{\beta'}(z|w,y=1,v=k)}{\sum_{j=1}^{K} p_{\beta'}(z|w,y=1,v=j) + p_{\beta^*}(z|w,y=1,v=K+1)} - \frac{p_{\beta'}(z|w,y=1,v=k)}{\sum_{j=1}^{K} p_{\beta'}(z|w,y=1,v=j)} \right| \\ &= \frac{p_{\beta^*}(z|w,y=1,v=K+1)p_{\beta'}(z|w,y=1,v=k)}{\left(\sum_{j=1}^{K} p_{\beta'}(z|w,y=1,v=j) + p_{\beta^*}(z|w,y=1,v=K+1)\right)} \\ &\times \frac{1}{\sum_{j=1}^{K} p_{\beta'}(z|w,y=1,v=k)} \\ &\leq \delta \frac{p_{\beta'}(z|w,y=1,v=k)}{\left(\sum_{j=1}^{K} p_{\beta'}(z|w,y=1,v=j)\right)^2} \\ &= \delta A(z,w,v=k) \,. \end{aligned}$$
(13)

Now we calculate ϵ_K given by

$$\mathbb{E}_{\mathcal{X}}[-\mathcal{L}(K;\phi'_{z},\phi'_{w},\beta',\theta')] - \mathbb{E}_{\mathcal{X}}[-\mathcal{L}(K+1;\phi^{*}_{z},\phi^{*}_{w},\beta^{*},\theta^{*})] = \epsilon_{K}$$

Because of (8), ϵ_K simplifies to

$$\begin{split} \epsilon_{K} &= \\ &- \mathbb{E}_{\mathcal{X}} \left[\mathbb{E}_{q_{\phi_{w}^{*}}(w|x,y=1)q_{\phi_{z}^{*}}(z|x)} \left[\sum_{j=1}^{K} p_{\beta'}(v=j|z,w,y=1) \log p_{\beta'}(z|w,y=1,v=j) \right] \right] \\ &+ \mathbb{E}_{\mathcal{X}} \left[\mathbb{E}_{q_{\phi_{w}^{*}}(w|x,y=1)q_{\phi_{z}^{*}}(z|x)} \left[\sum_{j=1}^{K+1} p_{\beta^{*}}(v=j|z,w,y=1) \log p_{\beta^{*}}(z|w,y=1,v=j) \right] \right] \\ &+ \mathbb{E}_{\mathcal{X}} \left[\mathbb{E}_{q_{\phi_{w}^{*}}(w|x,y=1)q_{\phi_{z}^{*}}(z|x)} \left[KL(p_{\beta'}(v|z,w,y=1)||p_{K}(v|y=1)) \right] \right] \\ &- \mathbb{E}_{\mathcal{X}} \left[\mathbb{E}_{q_{\phi_{w}^{*}}(w|x,y=1)q_{\phi_{z}^{*}}(z|x)} \left[KL(p_{\beta^{*}}(v|z,w,y=1)||p_{K+1}(v|y=1)) \right] \right] \\ &= \epsilon_{K}^{(1)} + \epsilon_{K}^{(2)} \end{split}$$

where $p_K(v|y=1)$ indicates that v is K-dimensional, and $\epsilon_K^{(1)}$ are the first two terms while $\epsilon_K^{(2)}$ are the the last two terms.

We first analyze $\epsilon_K^{(1)}$. For brevity, we combine the expectations and simply write $\mathbb{E}[\cdot]$. Together with (9), (11), and (13), we get

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$$\begin{split} \epsilon_{K}^{(1)} &|= \left| -\mathbb{E} \left[\sum_{j=1}^{K} p_{\beta'}(v=j|z,w,y=1) \log p_{\beta'}(z|w,y=1,v=j) \right] \right. \\ &+ \mathbb{E} \left[\sum_{j=1}^{K} p_{\beta^{*}}(v=j|z,w,y=1) \log p_{\beta'}(z|w,y=1,v=j) \right] \\ &+ \mathbb{E} \left[p_{\beta^{*}}(v=K+1|z,w,y=1) \log p_{\beta^{*}}(z|w,y=1,v=K+1) \right] \right| \\ &= \left| \mathbb{E} \left[\sum_{j=1}^{K} \log p_{\beta'}(z|w,y=1,v=j) \left(p_{\beta^{*}}(v=j|z,w,y=1) - p_{\beta'}(v=j|z,w,y=1) \right) \right] \right. \\ &+ \mathbb{E} \left[\frac{p_{\beta^{*}}(z|w,y=1,v=K+1) \log p_{\beta^{*}}(z|w,y=1,v=K+1)}{\sum_{j=1}^{K} p_{\beta'}(z|w,y=1,v=j) + p_{\beta^{*}}(z|w,y=1,v=K+1)} \right] \right| \\ &\leq \delta \cdot \mathbb{E} \left[\sum_{j=1}^{K} \left| \log p_{\beta'}(z|w,y=1,v=j) \right| A(z,w,v=j) \right] \\ &+ \left| \delta (\log \delta) \right| \mathbb{E} \left[\frac{1}{\sum_{j=1}^{K} p_{\beta'}(z|w,y=1,v=j)} \right] = o(1) \,, \end{split}$$

where the last inequality follows from $|x \log x|$ being increasing for $x \le 1/e$ and in o(1) we consider $\delta \to 0$.

Next we study $\epsilon_K^{(2)}.$ For shorthand, let us define

$$\log \left((K+1)p_{\beta^*}(v=K+1|z,w,y=1) \right)$$

=
$$\log \left(\frac{(K+1)p_{\beta^*}(z|w,y=1,v=K+1)}{\sum_{j=1}^{K} p_{\beta'}(z|w,y=1,v=j) + p_{\beta^*}(z|w,y=1,v=K+1)} \right)$$

=
$$\log p_{\beta^*}(z|w,y=1,v=K+1) + B(z,w)$$

and note that

$$\begin{split} |B(z,w)| &= \bigg| \log \left(\frac{(K+1)}{\sum_{j=1}^{K} p_{\beta'}(z|w,y=1,v=j) + p_{\beta^*}(z|w,y=1,v=K+1)} \right) \\ &\leq \max \left\{ \bigg| \log \left(\frac{(K+1)}{\sum_{j=1}^{K} p_{\beta'}(z|w,y=1,v=j)} \right) \bigg|, \\ & \bigg| \log \left(\frac{(K+1)}{\sum_{j=1}^{K} p_{\beta'}(z|w,y=1,v=j) + 1/e} \right) \bigg| \right\} \\ &= C(z,w). \end{split}$$

We have

$$\epsilon_{K}^{(2)} = \mathbb{E}\left[\sum_{j=1}^{K} p_{\beta'}(v=j|z,w,y=1) \log \left(Kp_{\beta'}(v=j|z,w,y=1)\right)\right]$$

$$-\mathbb{E}\left[\sum_{j=1}^{K} p_{\beta^{*}}(v=j|z,w,y=1)\log\left((K+1)(p_{\beta^{*}}(v=j|z,w,y=1))\right)\right] \\ -\mathbb{E}\left[p_{\beta^{*}}(v=K+1|z,w,y=1)\log\left((K+1)p_{\beta^{*}}(v=K+1|z,w,y=1))\right] \\ =\mathbb{E}\left[\sum_{j=1}^{K}(\log K)p_{\beta^{\prime}}(v=j|z,w,y=1) - (\log(K+1))p_{\beta^{*}}(v=j|z,w,y=1)\right] \\ +\mathbb{E}\left[\sum_{j=1}^{K}p_{\beta^{\prime}}(v=j|z,w,y=1)\log p_{\beta^{\prime}}(v=j|z,w,y=1) - p_{\beta^{*}}(v=j|z,w,y=1)\log p_{\beta^{*}}(v=j|z,w,y=1)\right] \\ -p_{\beta^{*}}(v=j|z,w,y=1)\log p_{\beta^{*}}(v=K+1|z,w,y=1)\right] \\ =\mathbb{E}\left[p_{\beta^{*}}(v=K+1|z,w,y=1)\log\left((K+1)p_{\beta^{*}}(v=K+1|z,w,y=1)\right)\right] \\ \geq \log(K) - (\log(K+1))\mathbb{E}\left[\sum_{j=1}^{K}p_{\beta^{*}}(v=j|z,w,y=1)\right] \\ +\mathbb{E}\left[\sum_{j=1}^{K}(p_{\beta^{\prime}}(v=j|z,w,y=1) - p_{\beta^{*}}(v=j|z,w,y=1))\log(p_{\beta^{\prime}}(v=j|z,w,y=1))\right] \right] \\ +\mathbb{E}\left[\sum_{j=1}^{K}p_{\beta^{\prime}}(z|w,y=1,v=K+1)\log p_{\beta^{*}}(z|w,y=1,v=K+1)\right] \\ -\left|\mathbb{E}\left[\left(\frac{p_{\beta^{*}}(z|w,y=1,v=j) + p_{\beta^{*}}(z|w,y=1,v=K+1)}{\sum_{j=1}^{K}p_{\beta^{\prime}}(z|w,y=1,v=j) + p_{\beta^{*}}(z|w,y=1,v=K+1)}\right)B(z,w)\right]\right] \\ \geq \log(K) - \log(K+1) \\ -\delta \cdot \mathbb{E}\left[\sum_{j=1}^{K}A(z,w,v=j)\left|\log(p_{\beta^{\prime}}(v=j|z,w,y=1))\right|\right]$$
(16)
 $-\delta (\log\delta) \mathbb{E}\left[\frac{1}{\sum_{j=1}^{K}p_{\beta^{\prime}}(z|w,y=1,v=j)}C(z,w)\right]$

$$\left\lfloor \sum_{j=1} p_{\beta'}(z|w, y=1, v=j) \right\rfloor$$
$$= \log \frac{K}{K+1} + o(1) .$$

In (15) we use (12), in (16) we rely on (13), and in (17) we use (14) again.

To summarize, we have $\epsilon_K \ge -|\epsilon_K^{(1)}| + \epsilon_K^{(2)} \ge -o(1) + o(1) + \log \frac{K}{K+1} = \log \frac{K}{K+1} + o(1)$. Thus $\epsilon_K \ge \log \frac{K}{K+1}$.